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Blind signal separation using an adaptive weibull distribution

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We propose an independent component analysis (ICA) algorithm which can separate mixtures of suband super- Gaussian source signals with self-adaptive nonlinearities. The ICA algorithem in the framework of natural Riemannian gradient is derived using the parameterized Weibull density model. The nonlinear function in ICA algorithem is self-adaptive and is controlled by the shape parameter of Weibull density model. Computer simulation results confirm the validity and high performance of the proposed algorithm

Key words: Independent component analysis, weibull distribution, maximum likelihood, sub- and super-Gaussian, blind signal separation.

INTRODUCTION

The problem of independent component analysis (ICA) has received wide attention in various fields such as biomedical signal analysis and processing (EEG, MEG and ECG), geophysical data processing, data mining, speech recognition, image recognition and wireless communications (Amari and Cichocki, 1998; Amari et al., 1997; Gardner, 1991; El-sayed Wahed and Mohamed, 2006). In many applications, the sensory signals (Observations obtained from multiple sensors) are generated by a linear generative model which is unknown to us. In other words, the observations are linear instantaneous mixtures of unknown source signals and the objective is to process the observations in such a way that the outputs correspond to the separate primary source signals. The operation starts with a random source vector S defined by $S(n) = [S_1, S_2, \dots, S_m]$ where the m components are supplied by a set of independent sources. Temporal sequences are considered here; henceforth the argument n denotes discrete time. The vector S is applied to a linear system whose input-output characterization is defined by a nonsigular m-by-m matrix A, called the mixing matrix. The result is an m-by-1 observation vector X (n) related to S (n) as follow X=AS where $X = [X_1, X_2, ..., X_m]^T$. The source vector S and the mixing matrix A are both unknown. The only thing available to us is the observation vector X. Given X, the problem is to find a demixing matrix W such that the original source vector S can be recovered from the output vector Y defined by Y = WX where $Y = [Y_1, Y_2, \dots, Y_m]^T$. This

is called the blind source separation. The solution to the blind source separation is feasible, except for an arbitrary scalling of each signal component and permutation of indices. In other words, it is possible to find a demixing matrix W whose individual rows are a rescaling and permutation of those of the matrix A, that is, the solution may be expressed in the form $Y=WX=WAS \rightarrow DPS$ where D is a nonsingular diagonal matrix and P is a permutation matrix.

Since Jutten and Herault (Karhunen, 1996) proposed a linear feedback network with a simple unsupervised learning algorithem, several methods have been developed.

Cichocki and Unbehauen (1996) and Comon (1994) proposed robust, flexible algorithm with equivariant properties. Comon (Diversi et al., 2005) gave a good insight to ICA problem from the statistical point of view. Bell and Seinowski (Bell and Seinowski, 1995) adopted an information maximization principle to find a solution to ICA problem. Maximum likelihood estimation (Alberg et al., 2002; Amari et al., 1997; El-sayed Wahed, 2007) was proposed by Pham et al. and was elaborated in (Abu-Taleb et al., 2006; El-sayed Wahed, 2007). The nonlinear extension of PCA was extensively studied in (Karhunen, 1996; El-sayed Wahed and Mohamed, 2006). Serial updating rule was introduced by Cardoso and Laheld (Cardoso and Laheld, 1996) and the resulting algorithm have equivariant performance. was shown to Independent, natural gradient was proposed and applied



Figure 1. Maximum Entropy Method.

to ICA by (Amari et al., 1996; Gardner, 1991; Gradshteyn et al., 1994). Conditions on cross cumulants for the separation of the source signals were investigated in (Alberg et al., 2002; Amari, 1998; Amari and Cichocki, 1998; Abu-Taleb et al., 2006; Choi et al., 1998; Choi and Cichocki, 1997).

Maximum entropy algorithm

This is an adaptive algorithm based on information theoretic approach and was suggested by Bell and Sejnowski (Bell and Sejnowski, 1995). The block diagram in Figure 1 explains the maximum entropy method for blind source separation.

The demixer operates on the observed data X to produce an output Y = WX, which is an estimate of source S. The output Y is transformed into Z by passing it through a non-linearity G (.), which is invertible and monotonic. For a given non-linearity G (.), the maximum entropy method produces an estimate of source S by maximizing the entropy h (Z) with respect to W. The mathematical representation of the whole process may be given as follows:

$$Z = G(y) = G (WAS) \qquad \Rightarrow S = A^{-1}W^{-1}G^{-1}(z) = \psi(z)$$

where G⁻¹ is the inverse non-linearity.

The probability density function of the output Z is defined in terms of that of the source S by

$$f[Z(z)] = \frac{f[S(s)]}{|\det(J(s))|} | s = \psi(z)$$

Where det (J(s)) is the determinant of the Jacobian matrix J(s). The *i*j-th element of the matrix J(s) is defined

$$J_{ij} = \frac{\partial Z_i}{\partial a}$$

by os_j . Hence, the entropy of the output Z at the output of the non-linearity G (.) is

$$\mathsf{h} (\mathsf{Z}) = -\mathsf{E}[\mathsf{log} f_z(\mathsf{Z})] = -\mathsf{E} \left[\mathsf{log} \left(\frac{f[\mathsf{S}(s)]}{|\mathsf{det}(\mathsf{J}(s))|} \middle| s = \psi(z) \right) \right] = -\mathsf{D}_{\mathrm{fs}} |\mathsf{det}(\mathsf{j})| \\ \text{evaluated } \mathsf{S} = \psi(z)$$

Hence, maximizing the entropy h (Z) is equivalent to minimizing the Kullback-Leibler divergence between $f_s(s)$

and a probability density function of S, defined by $|\det[J(s)]|$.

If the random variable z_i (ith element of z) is uniformly distributed inside the interval [0, 1] for all i, then the entropy h (z) is equal to zero. Accordingly,

h (Z) = -E[log
$$f_z(z)$$
] = -E $\left[log \left(\frac{f[S(s)]}{|det(J(s))|} \middle| s = \psi(z) \right) \right] \Rightarrow f_s(s) = |det(J(S))|$

Under the ideal condition, $W = A^{-1}$, the above relationship reduces to

$$\mathbf{f}_{\mathbf{S}_{i}}(\mathbf{S}_{i}) = \frac{\partial \mathbf{z}_{i}}{\partial \mathbf{y}_{i}} \Big|_{\mathbf{z}_{i}} = \mathbf{g}(\mathbf{s}_{i}) \text{ for all } \mathbf{i}.$$

Conversely, the results from Maximum Entropy Method may be stated as follows:

Let the non-linearity at the demixer output be defined in terms of the original source distribution as

$$z_i = g_i(y_i) = \int_{-\infty}^{z_i} fS_i(s_i) ds_i,$$

for all $i = 1, 2, ..., n.$

Then, maximizing the entropy of the random vector z at the output of the non-linearity G is equivalent to $W = A^{-1}$, which yields perfect blind source separation. The maximum entropy and maximum likelihood methods for blind source separation are equivalent under the condition that the random variable i z is uniformly distributed inside the interval [0, 1] for all i. This relationship may be proven with the help of chain rule of calculus as

$$\mathbf{J}_{ij} = \sum_{k=1}^{n} \frac{\partial z_i}{\partial z_i} \frac{\partial y_i}{\partial x_i} \frac{\partial x_i}{\partial s_i} = \sum_{k=1}^{n} \frac{\partial z_i}{\partial z_i} w_{ik} a_{kj}$$

The Jacobian matrix J is expressed as J = DWA, where D is a diagonal matrix given by

$$D = diag\left(\frac{\partial z_1}{\partial y_1}, \frac{\partial z_2}{\partial y_2}, \dots, \frac{\partial z_n}{\partial y_n}\right).$$

$$\left\| \det(\mathbf{J}) = \left\| \det(\mathbf{WA}) \right\|_{i=1}^{n} \frac{\partial g_{i}(\mathbf{y}_{i})}{\partial \mathbf{y}_{i}}$$

Hence

In the light of the above equation, an estimate of the probability density function $f_s(s)$ parameterized by the weight matrix W and the non-linearity G may be written formally as

$$f_{s}(s/W,G) = \left| det(WA) \right| \prod_{i=1}^{n} \frac{\partial g_{i}(y_{i})}{\partial y_{i}}$$
(1)

Therefore, under the above condition, maximizing the log-likelihood function {logfs (s/W, G)} {} is equivalent to maximizing the entropy h (Z) for blind source separation. Ell $f_{-}(z) = -E$

Referring to the expression h (Z) = -E[log f_z(Z)]

$$\begin{bmatrix} 1 & -E \\ 1 & -E \end{bmatrix} \begin{bmatrix} 1 & -E \\ 1 & -E \end{bmatrix}$$

$$\left\lfloor \log \left(\frac{1(S(3))}{|\det(J(s))|} \right| s = \psi(z) \right) \right\rfloor$$
, it is
source distribution is fixed, maximiz

seen that since the zing the entropy h(Z) requires maximizing the expectation of the denominator term $\{\log |\det(J(s))|\}$ with respect to the separating matrix W.

To do the computation using an adaptive algorithm that will maximize the objective function, the instantaneous objective function ϕ may be considered as:

$$\varphi = \frac{\log |\det(J)|}{2}$$
(2)

On expanding (2), we get:

$$\varphi = \log \left| \det(A) + \log \left| \det(W) \right| \right| + \sum_{i=1}^{n} \log \left(\frac{\partial z_i}{\partial y_i} \right)$$
 and

$$\frac{\partial \varphi}{\partial W} = W^{-T} + \sum_{i=1}^{n} \frac{\partial}{\partial W} \log \left(\frac{\partial z_i}{\partial y_i} \right)$$
(3)

The non-linear function should be judiciously selected to deal with the super-Gaussian, sub-Gaussian, stationary and non-stationary signals. The popular non-linearity used is logistic function and hyperbolic tangent function:

$$z_i = g(y_i) = \frac{1}{1 + e^{-y_i}},$$

 $z_i = g(y_i) = \tanh(y_i), \quad i = 1, 2, ..., n$

The non-linear functions should be monotonic and invertible.

 $\partial \varphi$

Finding out $\overline{\partial W}$ using the above non-linearity, we obtain 20

$$\frac{\partial \varphi}{\partial W} = W^{-T} + (1 - 2z)x^{T}$$

where x is the observed source vector, z is the non-linearly transformed output vector and 1 is a corresponding vector of ones.

Using the steepest ascent method to maximize the entropy h(Z), the change in weight matrix W is given by

$$\Delta W = \eta \frac{\partial \varphi}{\partial W} = \eta \left(W^{-T} + (1 - 2z)x^{T} \right)$$

where η is the learning rate parameter. The generalized final version for the update on W or the learning rule is obtained by using the natural gradient, which is equivalent to multiplying the expression for ΔW by W¹W instead of evaluating W^{T} as given below:

$$\Delta W = \eta \left(W^{-T} + (1-2z)x^{T} \right) W^{T}W = \eta \left(I + (1-2z)(Wx)^{T} \right) W = \eta \left(I + (1-2z)y^{T} \right) W$$
$$W(k+1) = W(k) + \eta \left(I + (1-2z(k))y^{T}(k) \right) W(k)$$
(4)

Where y is the output of the demixer before passing through the non-linearity, I is the unity matrix and is a fixed learning rate parameter with value less than 1.

The algorithm gives better result when applied on prewhitened data. It is sensitive to the learning rate parameter and works better for super-Gaussian signals.

WEIBULL MODEL FOR SOURCES

Optimal nonlinear activation function $f_s(s)$ is calculated by (1). However, it required the knowledge of the probability distribution of source signals which are not available to us. A variety of hypothesized density model has been used. For example, for the supper-Gaussian source signals, unimodal or hyperbolic-Cauchy distribution model (Bell and Sejnowski, 1995) leads to the nonlinear function given by

$$f_{s}(s) = \tanh(\beta f_{s}(s)).$$
(5)

Such sigmodal function was also used in (Bell and Sejnowski, 1995). For sub-Gaussian source signals, cubic nonlinear function $f_s(s) - f_s^3$ has been a favorite choice. For Mixtures of Sub- and super-Gaussian source signals, according to the estimated kurtosis of the expected signals, nonlinear function can be elected from two different choices (Diversi et al., 2005; Douglas et al., 1997) [For example, either $f_s(s) - f_s^3$ or $f_s(s) = tanh(\beta f_s(s))$]. several approaches (Girolami and Fyfe, 1997; Choi et al., 1998; Cichocki et al., 1997) are already available.

This paper present a flexible nonlinear function derived using Weibull density model. It will be shown that the



Figure 2. The plots of the standard Weibull density function for c = 0.25, 0.5, 1, 1.5.



Figure 3. The plots of the standard Weibull density function for c = 2, 3, 4, 5.

nonlinear function is self-adaptive and controled by Weibull shape parameter. It is not a form of fixed nonlinear function.

$$p(y;c,\zeta_0,\alpha) = \frac{c}{\alpha} \left(\frac{y-\zeta_0}{\alpha}\right)^{c-1} \exp\left(-\left(\frac{y-\zeta_0}{\alpha}\right)^c\right) \quad y \neq \zeta_0, \ c(\neq 0), \alpha \ (\neq 0)$$
(6)

The weibull distribution

The weibull probability distribution is a set of distributions parameterized by a positive real number c which is usually referred to as the shape parameter of the distribution. The shape parameter c controls the peakiness of the distribution. The probability density function (PDF) for Weibull is described by It is necessary c be greater than -1, for otherwise the integral of $p(y;c,\zeta_0,\alpha)$ between $y = \theta$ and $y = \hat{\theta} \phi \theta$ will be infinite. The standard form of the distribution will have $\zeta_0 = 0$ and $\alpha = 1$ so that the standard density function is $p(y;c,\zeta_0,\alpha) = \frac{c}{\alpha} y^{c-1} \exp(-y^c) \quad y \phi 0, \ c(\phi 0)$ (7)



Figure 4. The plot of kurtosis k(c) versus the shape parameter c for leptokurtic and platykurtic signals.

The distribution of y now depends on the shape parameter c alone. The plots of the standard density function in (7) for c = 0.25, 0.5, 1, 1.5, 2, 3, 4, 5 are presented in Figures 2 and 3 respectively.

The moments of weibull distribution

In order to fully understand the Weibull distribution, it is useful to look at its moments (specially 2^{nd} and 4^{th} moments which give the kurtosis). The n^{th} moment of Weibull distribution is given by:

$$E(y^{r}) = \int_{0}^{\infty} y^{r} p(y; c) dy = \Gamma\left(\frac{r}{c} + 1\right)$$
(8)

Then

$$M_{2} = \Gamma\left(\frac{2}{c} + 1\right)$$

$$M 4 = \Gamma\left(\frac{4}{c} + 1\right)$$

$$M_{2k} = \Gamma\left(\frac{2k}{c} + 1\right)$$
(9)

The moment ratios, coefficient of variation, and standard

$$\frac{k_r}{k^{r/2}}$$

cumulants κ_2 of the standard distribution in (7) are of course the same as those of the distribution in (6).

Kurtosis and shape parameter

The kurtosis is anondimensional quantity. It is measures the relative peakdness or faltness of a distribution. A

distribution with positive kurtosis is termed leptokurtic (super-Gaussian). A distribution with negative kurtosis is termed platykurtic (sub-Gaussian). The kurtosis of the distribution is defined in terms of the 2^{nd} -and 4^{th} –order moment as

$$k(y) = \frac{M_4}{M_2^2} - 3 \tag{10}$$

1

Where the constant term –3 makes the value zero for the standard normal distribution. For Weibull distribution, the kurtosis can be expressed in terms of the shape parameter, given by

$$k(c) = \frac{\Gamma\left(\frac{4}{c}+1\right)}{\left(\Gamma\left(\frac{2}{c}+1\right)\right)^2} - 3$$
(11)

The plots of kurtosis k(c) versus the shape parameter c for leptokurtic and platykurtic signals are shown in Figure 4. The activation function for Weibull distribution in (3) is given by

$$f_i(y_i) = \frac{c - 1 - cy^c}{y} \tag{12}$$

$$\frac{\partial L}{\partial c_i} = \frac{\partial \ln p(y_i; c_i)}{\partial c_i} = \frac{1}{c_i} + \ln(y_i) - y_i^{c_i} \ln(y)$$
(13)

$$\Delta c_i = -\eta_i \frac{\partial L}{\partial c_i} = -\eta_i (\frac{1}{c_i} + \ln(y_i) - y_i^{c_i} \ln(y))$$
(14)

COMPUTER SIMULATION RESULTS

Consider the system involving the following three independent sources

 $U_1(n)$ = a square wave of amplitude a, and fundamental frequency w0

 $U_2(n)$ = a triangular wave of amplitude, and fundamental frequency w0

$$U_3(n) = 0.1\sin(400n)\cos(30n)$$
(15)

The mixing matrix A is

$$\mathbf{A} = \begin{bmatrix} 0.56 & 0.79 & -0.37 \\ -0.75 & 0.65 & 0.86 \\ 0.17 & 0.32 & -0.48 \end{bmatrix}$$
(16)

The algorithm was implemented using the following conditions:



Figure 5. The original and demixed signals.

i.) Initialization. The weights in the demixing matrix W were picked from a random number generator with a uniform distribution inside the range [0.0, 0.5].

ii.) The learning rate parameter was fixed at $\eta = 0.1$

iii.) Signal duration. The time series produced at the mixer output had a sampling period 10^{-4} s and contained N = 65,000 samples.

Figure (5) displays the waveforms of the source signals and the signals produced at the output of the demixer. It can be observed that after 3000 iterations, source signals are well separated.

Conclusion

We have proposed ICA algorithem (in the framework of natural Riemannian gradient) where the self-adaptive nonlinear function easy derived using Weibull density model for the probability distributions of the source signals. We have shown that the proposed ICA algorithm can separate the mixtures of sub-and super-gaussian signals with self adaptive nonlinearities which is controlled by Weibull shape parameter. Finally we apply our algorithm on a mixture of other distributions and image separation, which give a good result.

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